MULTIPLE IMPUTATION



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STRATEGIES FOR HANDLING MISSING DATA

- Complete Case Analysis (Available Case Analysis)
- Single Imputation

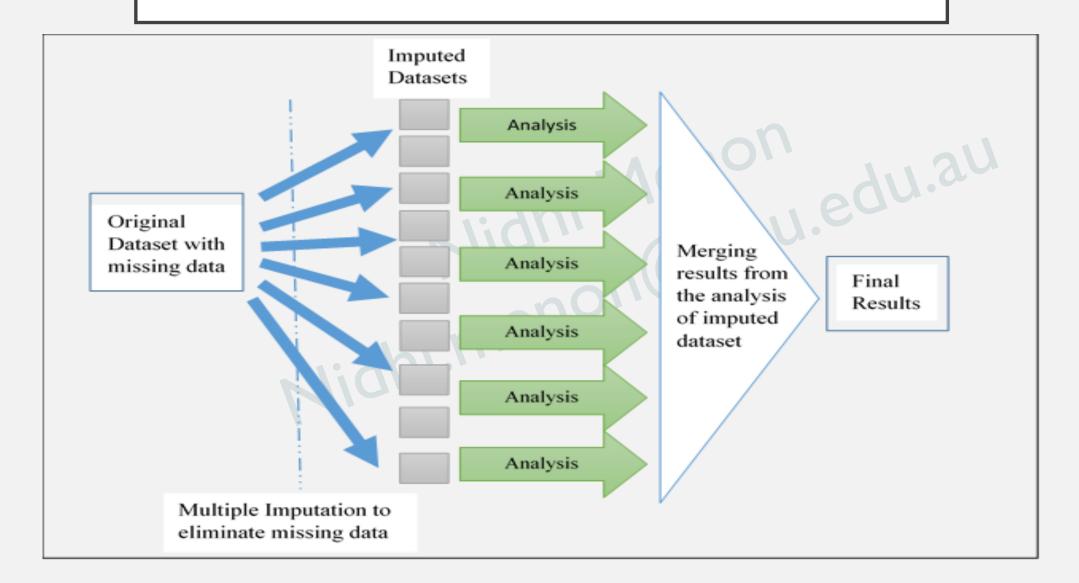


DEVELOPMENT OF MULTIPLE IMPUTATION

- **1987:** Inception Donald. B. Rubin
- 1987: 1st edition of Statistical Analysis with Missing Data by Little and Rubin
- **1997:** NORM, Schafer
- **1999:** MICE- concept, Van Buuren
- **2002:** SAS implemented the mi routine

- 2008: mice in R, Van Buuren
- 2011: Amelia was released in R
- 2012: MI in Stata
- 2016:mice in multilevel data, Ian White
- **2017:** jomo in R, Carpenter & Quartagno
- **2018:** micemd, Vinvent Audigier

MULTIPLE IMPUTATION



Mechanism of Missingness

Mechanism

Structure of Imputation Model



Selecting Predictors for the Imputation Model

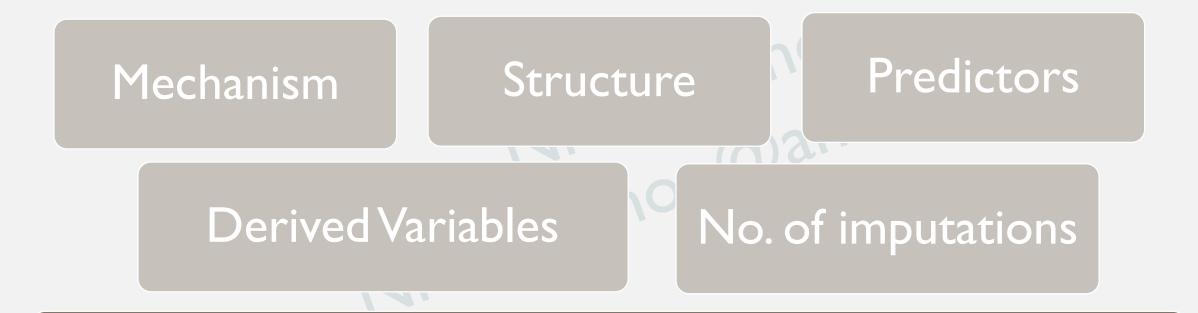




Imputing Derived Variables



Number of Imputations



Order of Imputations

RECIPE FOR IMPUTATION

- 1. Define the most general analytic model to be applied to imputed data
- 2. The target variable is the variable with the missing values
- 3. Select a method that imputes close to the data
- 4. Include all level-I variables and their cluster means
- 5. Include all level-2 predictors
- 6. Include any interactions implied by the model
- 7. Exclude any terms involving the target variable



METHODS OF IMPUTATION – FCS/ MICE

- Can be used for datasets containing both continuous and categorical data.
- Defines an imputation model on a variable by variable basis -> great for datasets with complex structures
- The method also allows the researcher to account for the complexities observed in the data, in the imputation model.
- Consider a scenario with 3 partially missing covariates namely X₁,X₂ and X₃ and outcome variable Y is complete. Here, X₁ = [X₁^{mis};X₁^{obs}]; X₂ = [X₂^{mis};X₂^{obs}] & X₃=[X₃^{mis};X₃^{obs}]

```
\begin{array}{l} \textit{Iteration (1):} \\ & \theta_1^{(1)} \sim f(\theta_1).f(x_1^{obs}|x_2^{obs}.\,x_3^{obs}.\,y.\,\theta_1) \\ & x_1^{mis(1)} \sim f(x_1^{mis}|\,x_2^{obs}.\,x_3^{obs}.\,y.\,\theta_1^{(1)}) \\ & \theta_2^{(1)} \sim f(\theta_2).f(x_2^{obs}|x_1^{obs(1)}.\,x_3^{obs}.\,y.\,\theta_2) \\ & x_2^{mis(1)} \sim f(x_2^{mis}|\,x_1^{obs(1)}.\,x_3^{obs}.\,y.\,\theta_2^{(1)}) \\ & \theta_3^{(1)} \sim f(\theta_3).f(x_3^{obs}|x_1^{obs(1)}.\,x_2^{obs(1)}.\,y.\,\theta_3) \\ & x_3^{mis(1)} \sim f(x_3^{mis}|\,x_1^{obs(1)}.\,x_2^{obs(1)}.\,y.\,\theta_3^{(1)}) \end{array}
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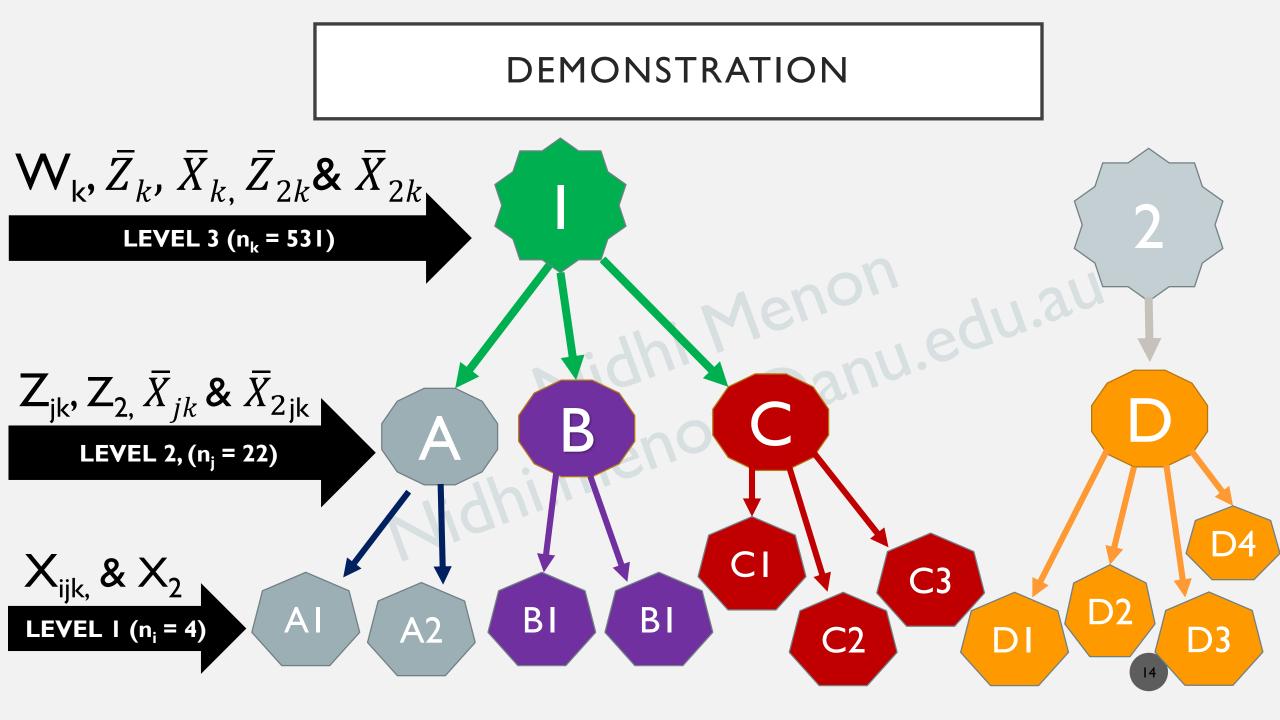
METHODS OF IMPUTATION – JOMO

- Defines a multivariate joint model for all variables in the dataset for imputation of missing values in the outcome.
- Outcome here refers to the variables in the model with missing values and not the outcome of the analysis model.
- Suppose we have variables Y_1 and Y_2 are partially observed and Y_3 and Y_4 are variables with no missing data, then the simplest joint model is the multivariate normal model given by:

 $Y_{1,i} = \beta_{01} + \beta_{11} Y_{3,i} + \beta_{21} Y_{4,i} + e_{1,i}$ $Y_{2,i} = \beta_{01} + \beta_{11} Y_{3,i} + \beta_{21} Y_{4,i} + e_{2,i}$ $\binom{e_{1,i}}{e_{2,i}} \sim N(0, \Omega)$

• Jomo used the Gibbs Sampling approach by consistently drawing new values for all parameters i.e. the fixed effects (β), the covariance matrix and the missing data.

• The current draw of missing values is combined with the observed data to make the first imputed dataset



ANALYSIS MODEL

$$Y_{ijk} = \beta_{0jk} + \beta_{1jk} (X_{ijk} - \bar{X}_{jk}) + e_{ijk}; \ e_{ijk} \sim N(0, \sigma^2)$$

$$\beta_{0jk} = \gamma_{00k} + \gamma_{01} (Z_{jk} - \bar{Z}_k) + \gamma_{02} (\bar{X}_{jk} - \bar{X}_k) + r_{0jk}; \ r_{0jk} \sim N(0, \tau_r)$$

$$\beta_{1j} = \gamma_{10}$$

$$\gamma_{00k} = \delta_{000} + \delta_{001} (W_k) + u_{00k}; \ u_{00k} \sim N(0, \tau_u)$$

Substituting, we get

 $Y_{ijk} = \delta_{000} + \gamma_{10}(X_{ijk} - \bar{X}_{jk}) + \gamma_{02}(\bar{X}_{jk} - \bar{X}_k) + \gamma_{01}(Z_{jk} - \bar{Z}_k) + \delta_{001}W_k + u_{00k} + r_{0jk} + e_{ijk}W_k + v_{00k} + v_{$

Here, $i = 1, 2, ..., n_{jk}, j = 1, 2, ..., n_k \& k = 1, 2, ..., K$.

DEMONSTRATION

e_{ijk} r_{0jk} u_{00k}	$ \begin{array}{l} rnorm(n = i * j * k, mean = 0, sd = 1) \\ rep(rnorm(n = k * j, mean = 0, sd = 1), each = i) \\ rep(rnorm(n = k, mean = 0, sd = 1), each = i * j) \end{array} $	
Coefficients	$g_{000} = 2, \ g_{100} = 2.5, \ g_{200} = 2.5, \ g_{010} = 2, g_{001} = 3$	
Y_{ijk}	$\delta_{000} + \gamma_{10}(X_{ijk} - \bar{X}_{jk}) + \gamma_{02}(\bar{X}_{jk} - \bar{X}_k) + \gamma_{01}(Z_{jk} - \bar{Z}_k) + \delta_{001}W_k + u_{00k} + r_{0jk} + e_{ijk}$	
X_{ijk}	rsn(n = i * j * k, xi = 70, omega = 20, alpha = 10)	
Z_{jk} W_k	rep(rtpois(j * k, lambda = 3, a = 2, b = 25), each = i) rep(rtpois(k, lambda = 3, a = 2, b = 10), each = i * j)	
Correlated Variables:		
$Z X_2$	rtruncnorm(n = i * j * k, a = 59, b = 396, mean = 114.8, sd = 14) (0.6) * X _i jk + sqrt(1 - 0.6) * Z; correlation = 0.63	
Z_1 Z_2	rep(rtpois(j * k, lambda = 5.8, a = 1, b = 41), each = i (0.87) * Zjk + sqrt(1 - 0.87) * Z1; correlation = 0.7	

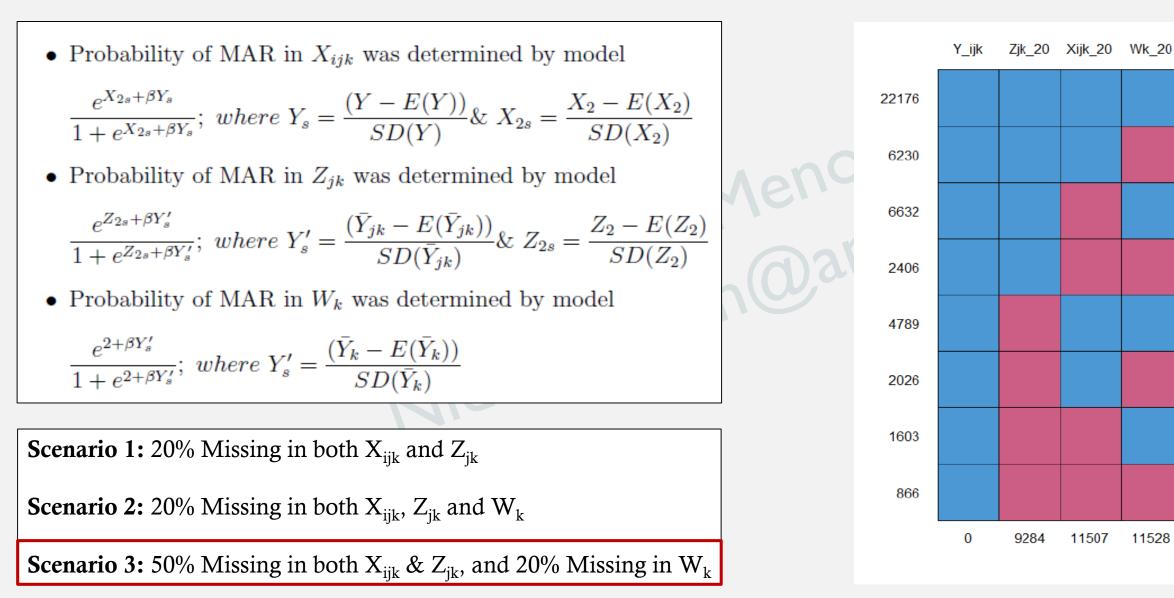
ANALYSIS ON ORIGINAL DATASET

Table 3: Analysis of the complete (simulated) dataset

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Fixed Effect	Coefficient	se	p-value	equ.
Intercept	1.851	0.618	< 0.001	
$X_{ijk} - \bar{X}_{jk}$	2.5	$4.385e^{-04}$	< 0.001	
$\begin{array}{l} X_{ijk} - \bar{X}_{jk} \\ \bar{X}_{jk} - \bar{X}_k \end{array}$	2.498	$1.767e^{-03}$	< 0.001	
$Z_{jk} - \bar{Z}_k$	2.001	$8.341e^{-02}$	< 0.001	
W_k	3.070	$3.712e^{-02}$	< 0.001	
Random Effects	Variance	Std. Dev		-
Level 3 effect (u_{00k})	0.9969	0.9984		-
Level 2 effect (r_{0jk})	1.0354	1.0176		
Level 1 effects (e_{ijk})	0.9996	0.9998		

INTRODUCING MISSING DATA

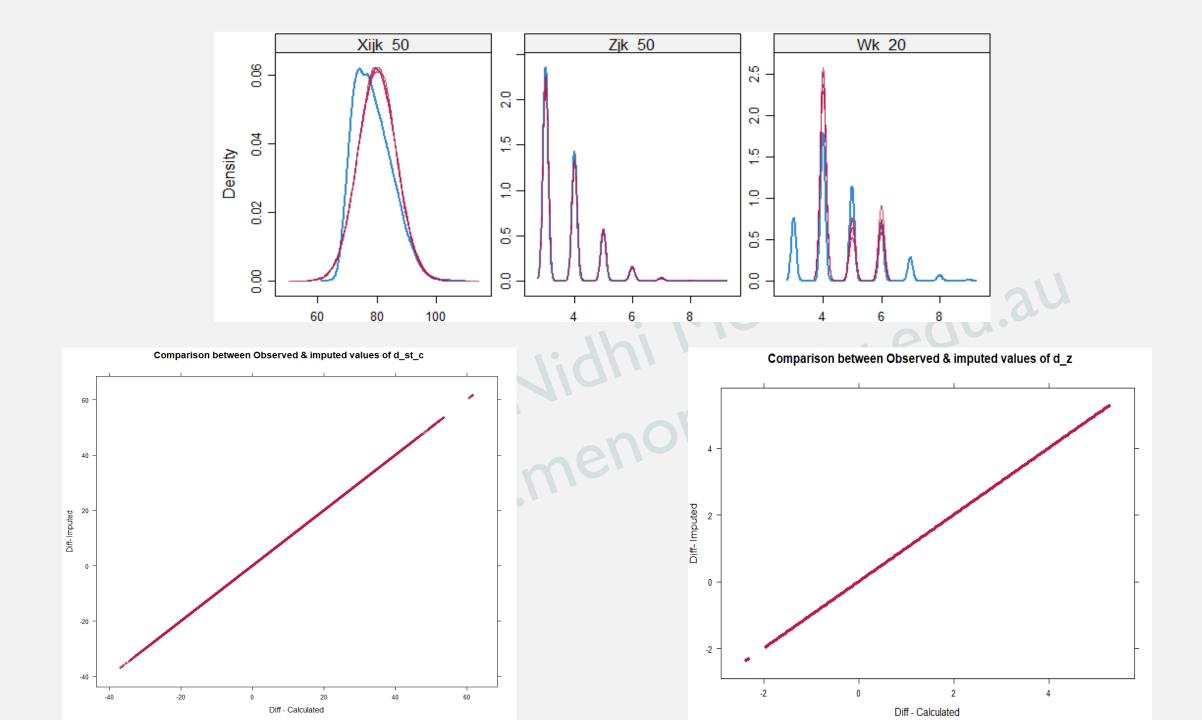
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Available Case Analysis

Scenario - 3	(n = 11,083)		
Fixed Effect	Coefficient	se	p-value
Intercept	-218	0.703	< 0.001
$X_{ijk} - X_{jk}$	2.495	$1.801e^{-03}$	< 0.001
$\bar{X}_{jk} - \bar{X}_k$	2.597	$8.866e^{-02}$	< 0.001
$Z_{jk} - \bar{Z}_k$	2.024	$2.159e^{-02}$	< 0.001
$\tilde{W_k}$	2.784	0.143	< 0.001
Random Effects	Variance	Std. Dev	
Level 3 effect (u_{00k})	1.012	1.006	
Level 2 effect (r_{0jk})	11.701	3.421	
Level 1 effects (e_{ijk})	1.010	1.005	

Random Effects	Variance	Std. Dev					
Level 3 effect (u_{00k})	1.012	1.006					
Level 2 effect (r_{0jk})	11.701	3.421					
Level 1 effects (e_{ijk})	1.010	1.005					
			Nid			MICE	
Scenario - 2	JOMO			Scenario - 3			
	<i>a m i i</i>				a t a a		
Fixed Effect	Coefficient	se	p-value	Fixed Effect	Coefficient	se	p-value
Intercept	1.655	0.178	< 0.001	Intercept	2.877	0.7640	< 0.001
$\overline{X_{ijk}} - \overline{X}_{jk}$	1.775	0.008	< 0.001	$X_{ijk} - \overline{X}_{jk}$	1.876	0.0396	< 0.001
$\bar{X}_{jk} - \bar{X}_k$	1.7637	0.0109	< 0.001	$\bar{X}_{jk} - \bar{X}_k$	2.407	0.0432	< 0.001
$Z_{jk} - \bar{Z}_k$	1.763	0.0781	< 0.001	$Z_{jk} - \bar{Z}_k$	2.9681	0.2130	< 0.001
W_k	3.0701	0.0404	< 0.001	W_k	2.5492	0.1730	< 0.001
Random Effects	Variance			Random Effects	Variance	Std. Dev	
Level 3 effect (u_{00k})	2.5263			Level 3 effect (u_{00k})	$2.0894e^{-14}$		
Level 2 effect (r_{0ik})	1.3098e - 13			Level 2 effect (r_{0jk})	0.7294		
Level 1 effects (e_{ijk})	1.0943			Level 1 effects (e_{ijk})	7.622		





DISADVANTAGES

Disadvantages of JoMo:- Considering a joint model on variables subject to missingness may not always be feasible or even realistic. For example, consider a survey with items targeted at different sub-populations; e.g. item asking respondents when was their last pap smear or item asking respondents the number of cigarettes smoked in the last 24 hours. This could apply to even questionnaires with a skip pattern. Imposing a joint distribution when a joint distribution may not even exist is not practicable. There are several cases when a joint modelling strategy may not work such as when variables have nominal, count or semi-continuous variables (Yucel, 2008). Thus researchers must remain cautious when choosing the right method of imputation bearing these factors in mind.



Disadvantages of FCS:- Although an appealing method of imputation, FCS is not without its limitations. The method is based on the assumption that the data is missing at random (MAR). Secondly, each conditional distribution needs to be specified separately. This would result in substantial modelling especially for datasets with many variables. The technique is more computationally challenging compared to joint modelling.