



Interrupted time series

Alice Richardson

National Centre for Epidemiology & Population Health

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Overview

- Introduction
- Examples
- Stata
- R
- Extensions



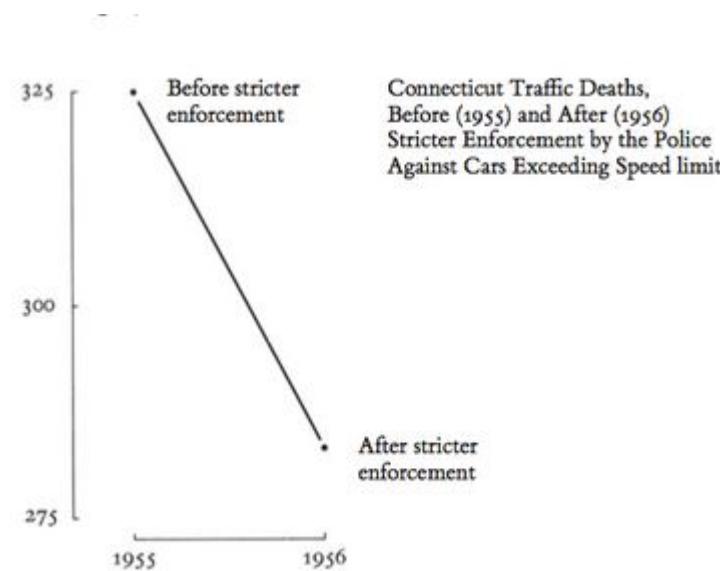
Why use an interrupted time series?

- A randomised experiment is not practical but a quasi-experiment is
- Data collected at equally-spaced time intervals
- Series “interrupted” by an intervention
- Measure effect of the intervention
- Control for time trends



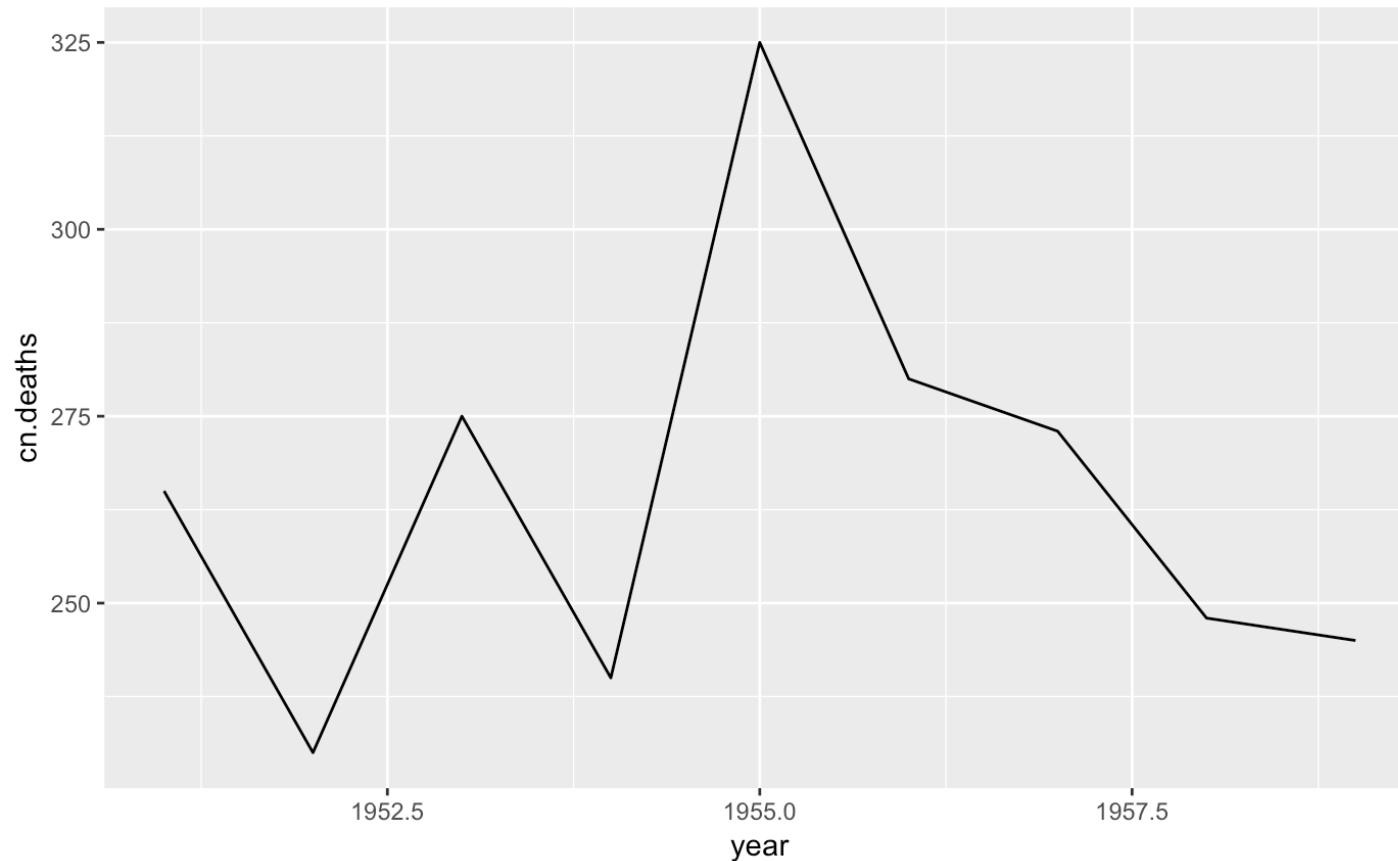
Examples

- Speeding in Connecticut: did the introduction of stricter enforcement reduce fatalities?



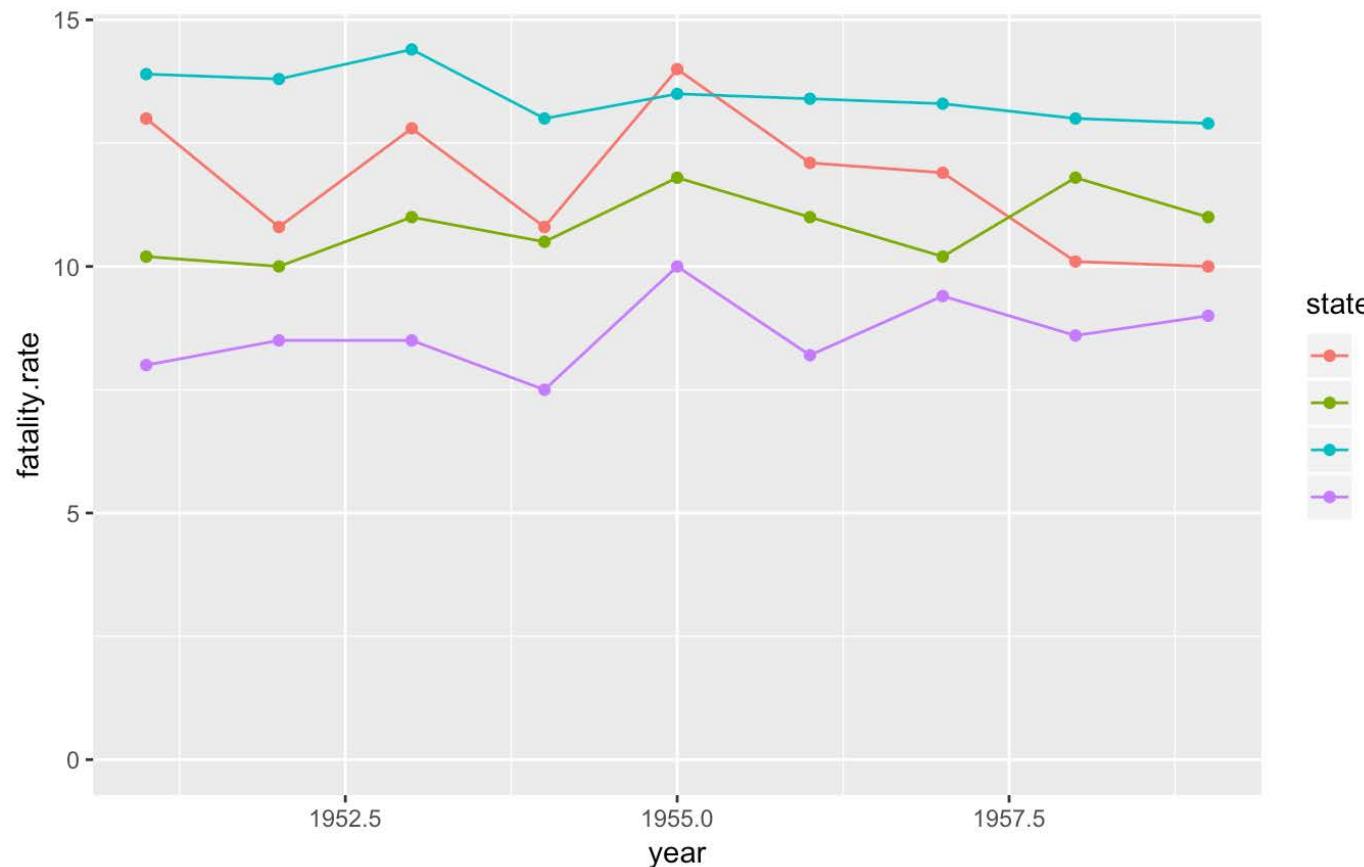


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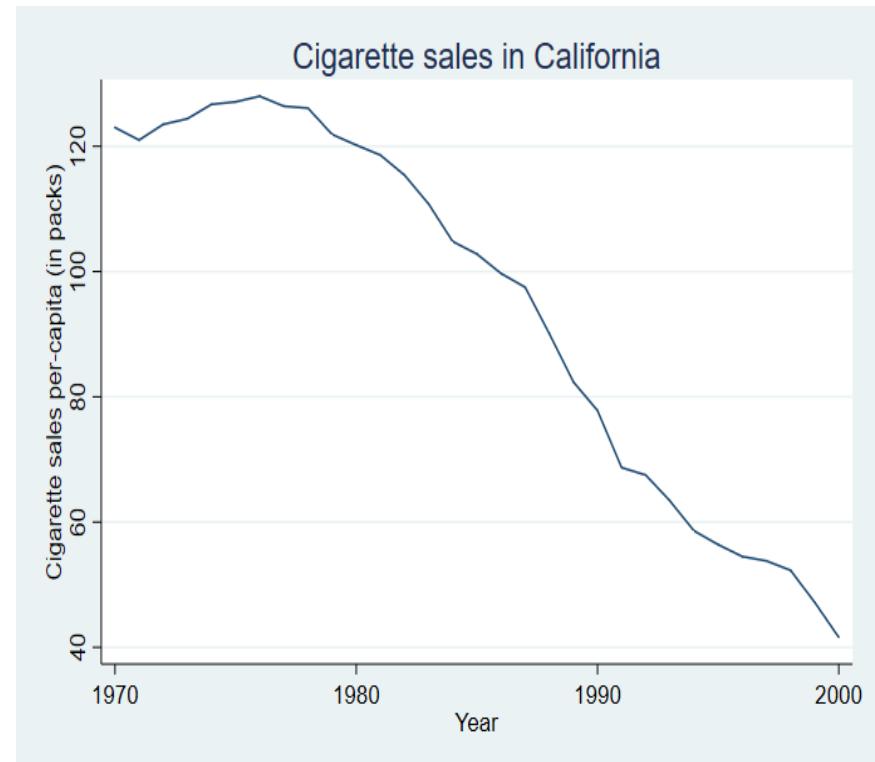


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- Cigarette sales: did a tax increase beginning in 1989 reduce sales?

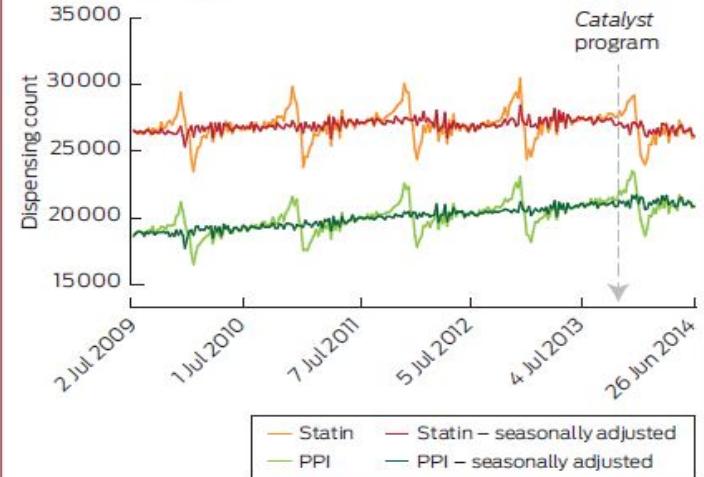




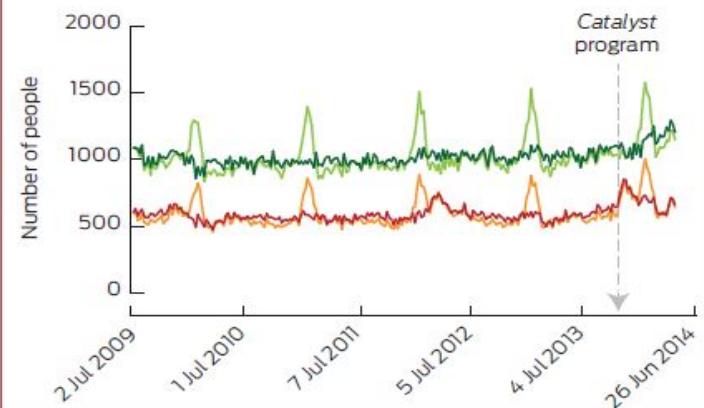
- Catalyst and statin sales: did the TV show influence people to stop taking their statins?

1 Weekly unadjusted and seasonally adjusted (A) dispensing counts and (B) number of people discontinuing use of statins and proton pump inhibitors (PPIs), 1 July 2009 to 30 June 2014

A: Dispensing counts

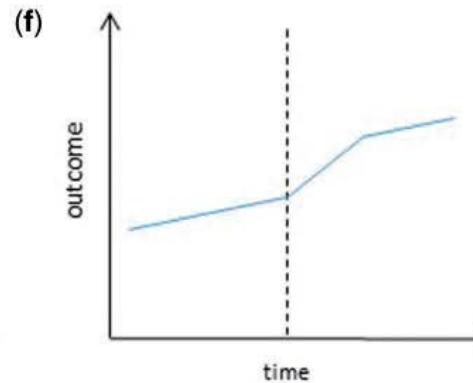
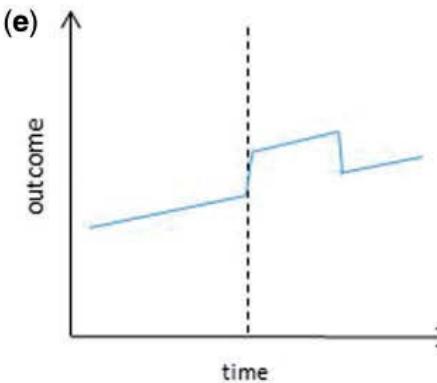
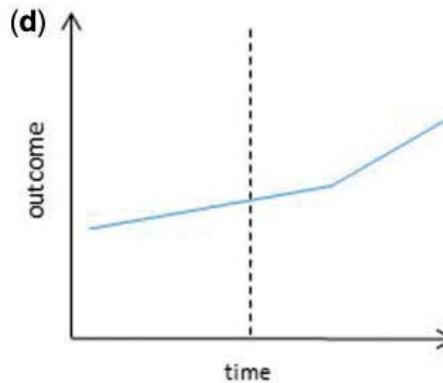
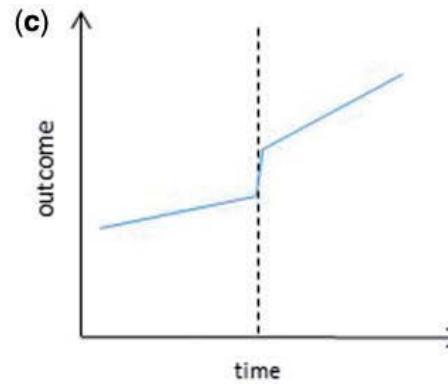
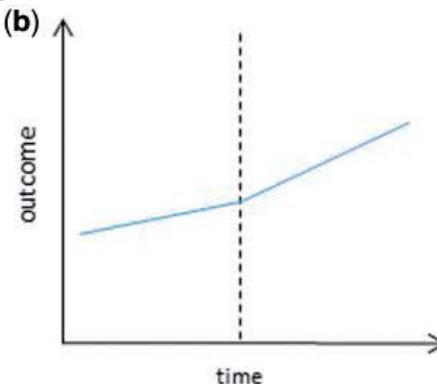
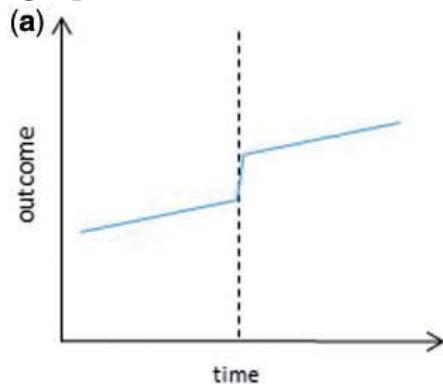


B: Number of people discontinuing use





Types of impact



- Level change; Slope change; Level and slope change; Slope change with a lag; Temporary level change; Temporary slope change leading to a level change

Stata: set up and fit model

```
. use http://fmwww.bc.edu/repec/bocode/c/cigsales.dta

.

.

.

. tsset state year
    panel variable: state (strongly balanced)
    time variable: year, 1970 to 2000
        delta: 1 unit

.

.

.

. itsa cigsale, single treat(3) trperiod(1989) lag(1) posttrend figure

    panel variable: state (strongly balanced)
    time variable: year, 1970 to 2000
        delta: 1 unit
```



Regression with Newey-West standard errors
maximum lag: 1

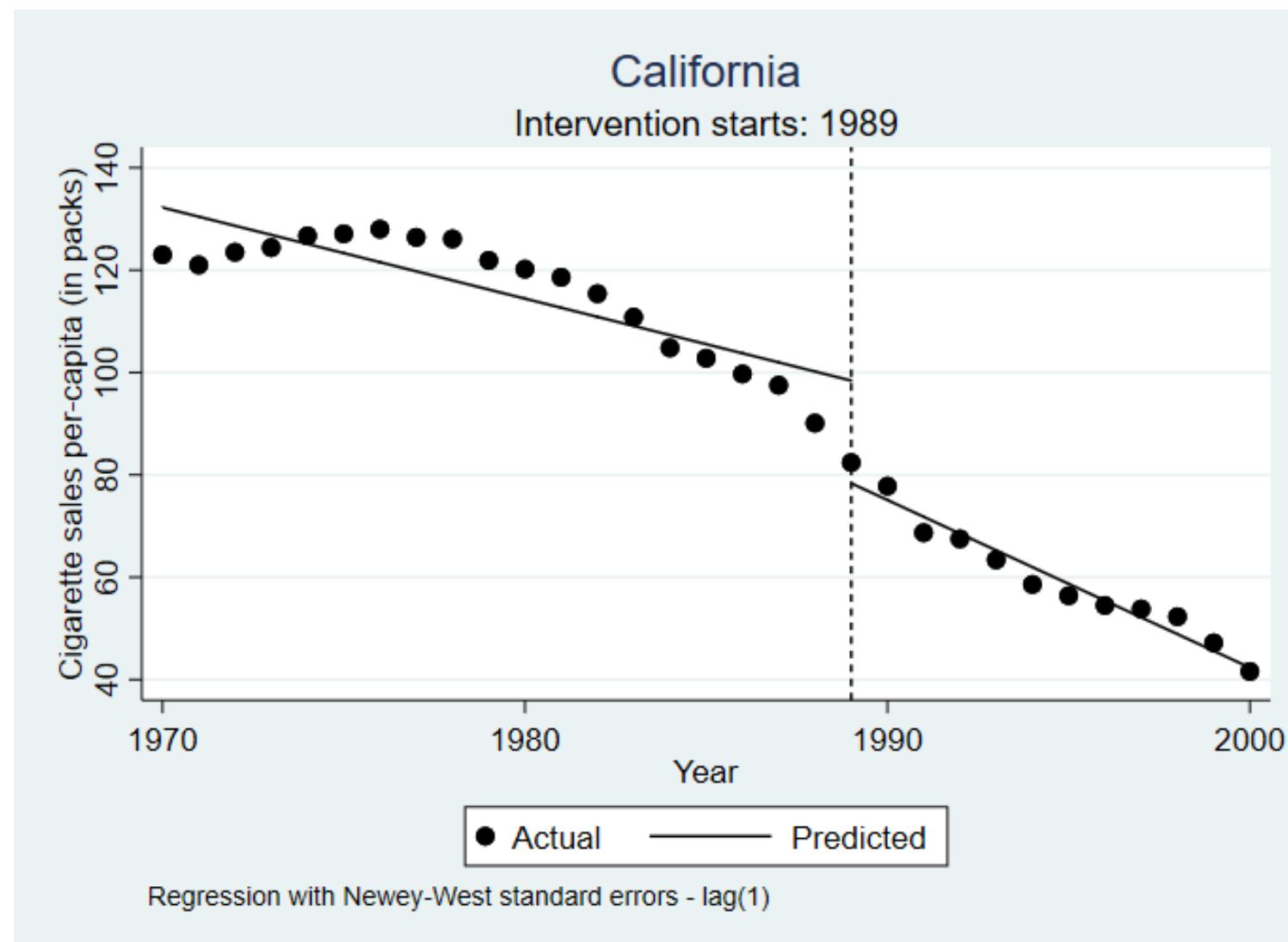
Number of obs	=	31
F(3, 27)	=	331.45
Prob > F	=	0.0000

cigsale	Newey-West					
	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
_t	-1.779474	.3834188	-4.64	0.000	-2.566184	-.9927632
_x1989	-20.0581	4.724395	-4.25	0.000	-29.75175	-10.36444
_x_t1989	-1.494652	.4368201	-3.42	0.002	-2.390933	-.5983715
_cons	132.2258	4.253054	31.09	0.000	123.4992	140.9523

Postintervention Linear Trend: 1989

Treated: $_b[_t]+_b[_x_t1989]$

Linear Trend	Coeff	Std. Err.	t	P> t	[95% Conf. Interval]	
Treated	-3.2741	0.2688	-12.1803	0.0000	-3.8257	-2.7226





Test for autocorrelation

```
. acetest, lag(6)
```

Cumby-Huizinga test for autocorrelation

H0: variable is MA process up to order q

HA: serial correlation present at specified lags >q

H0: q=0 (serially uncorrelated) HA: s.c. present at range specified				H0: q=specified lag-1 HA: s.c. present at lag specified			
lags	chi2	df	p-val	lag	chi2	df	p-val
1 - 1	15.242	1	0.0001	1	15.242	1	0.0001
1 - 2	15.255	2	0.0005	2	3.300	1	0.0693
1 - 3	15.325	3	0.0016	3	1.192	1	0.2749
1 - 4	15.896	4	0.0032	4	0.000	1	0.9880
1 - 5	16.057	5	0.0067	5	1.113	1	0.2914
1 - 6	16.078	6	0.0133	6	2.051	1	0.1521

Test allows predetermined regressors/instruments

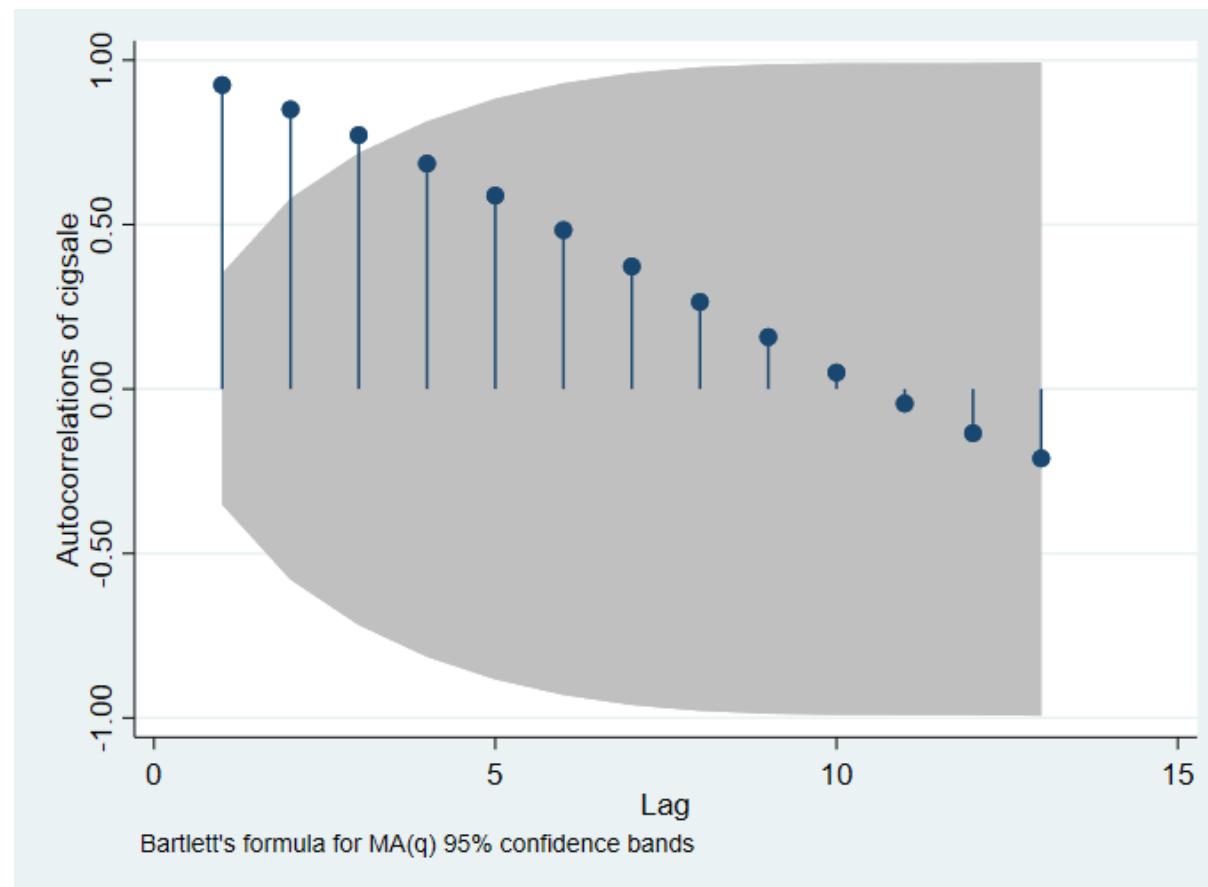
Test requires conditional homoskedasticity



Autocorrelation plots

```
. ac cigsale  
. corrrgram cigsale
```

LAG	AC	PAC	Q	Prob>Q	-1	0	1	-1	0	1
					[Autocorrelation]			[Partial Autocor]		
<hr/>										
1	0.9246	1.0380	29.154	0.0000	-----			-----		
2	0.8504	-0.4160	54.665	0.0000	-----			---		
3	0.7721	-0.2780	76.444	0.0000	-----			--		
4	0.6855	0.0421	94.248	0.0000	-----					
5	0.5881	0.1094	107.86	0.0000	----					
6	0.4834	-0.2862	117.42	0.0000	---			--		
7	0.3727	-0.0382	123.34	0.0000	--					
8	0.2649	0.2119	126.46	0.0000	--			-		
9	0.1582	0.0841	127.62	0.0000	-					
10	0.0504	-0.3709	127.75	0.0000				--		
11	-0.0441	-0.1241	127.85	0.0000						
12	-0.1343	-0.4970	128.82	0.0000	-			---		
13	-0.2109	-0.7565	131.35	0.0000	-			-----		





Alternatively, model AR(1) directly

- `itsa cigsale, single treat(3) trperiod(1989)
replace prais rhoype(tscorr) vce(robust)`

Prais-Winsten AR(1) regression -- iterated estimates

Linear regression

	Number of obs	=	31
	F(3, 27)	=	609.24
	Prob > F	=	0.0000
	R-squared	=	0.9011
	Root MSE	=	2.5964

cigsale	Semirobust					
	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
_t	-1.843139	.4538631	-4.06	0.000	-2.77439	-.9118892
_x1989	-6.094491	.8840197	-6.89	0.000	-7.90835	-4.280633
_x_t1989	-1.998494	.9191	-2.17	0.039	-3.884332	-.1126568
_cons	126.35	3.789489	33.34	0.000	118.5746	134.1254
rho	.9424635					

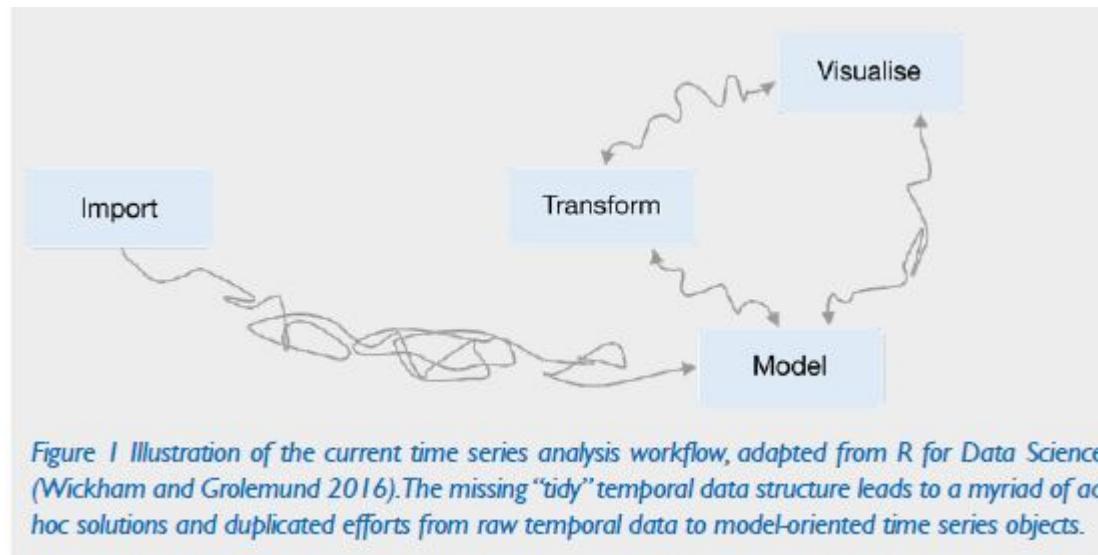
Durbin-Watson statistic (original) 0.535242

Durbin-Watson statistic (transformed) 1.342728



R set up and fit model

- `data$smokban <- c(rep(0, 19), rep(1, nrow(data) - 19))`
- `data$year1970 <- data$year - 1970`
- `cigsale.model01 <- lm(cigsale ~ smokban*year1970, data = data)`
- `summary(cigsale.model01)`

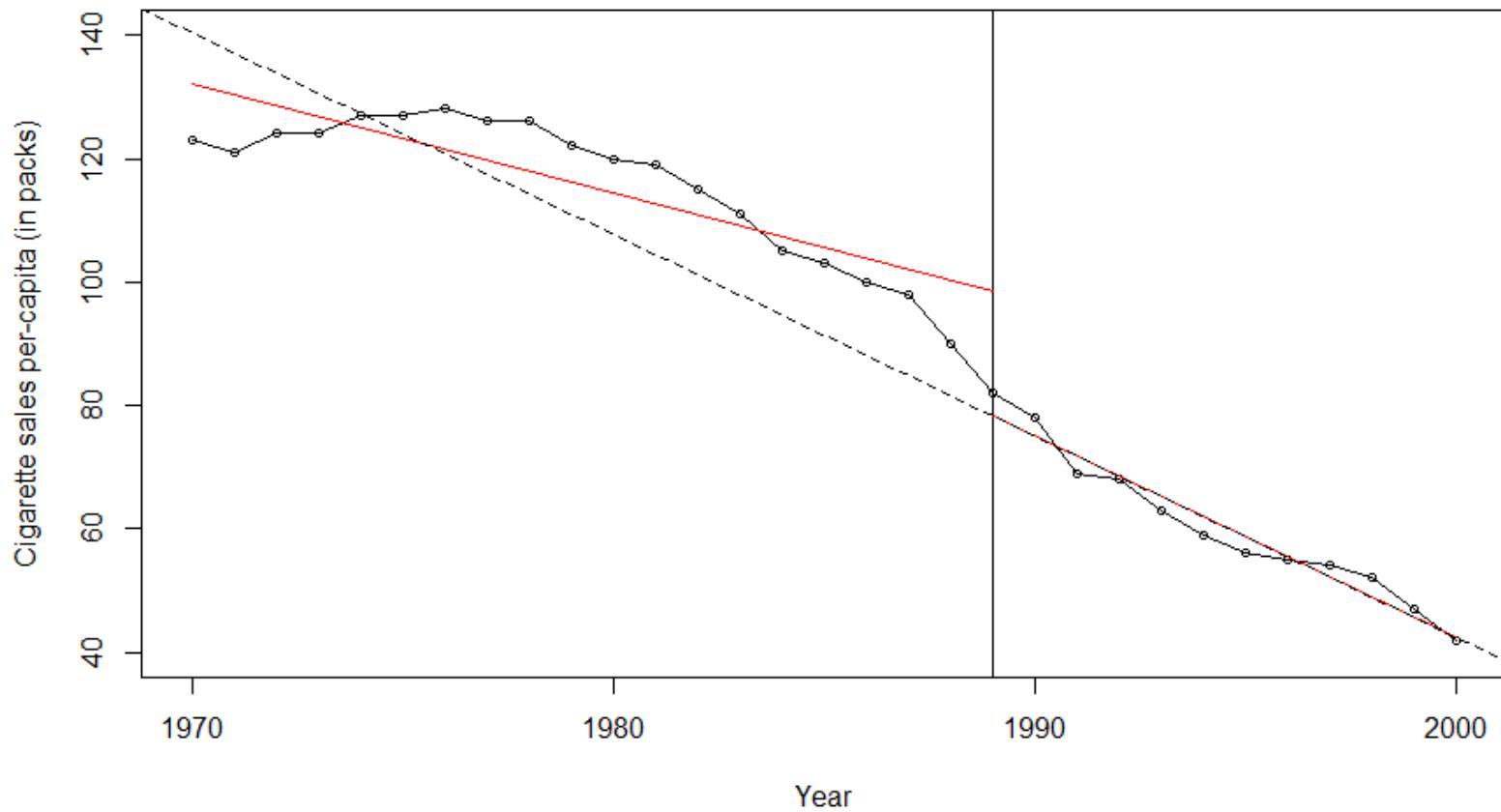


- Call:
- ```
lm(formula = cigsale ~ smokban * year1970, data = data)
```
- Coefficients:

|                  | Estimate | Std. Error | t value | Pr(> t )     |
|------------------|----------|------------|---------|--------------|
| (Intercept)      | 132.1789 | 2.2528     | 58.673  | < 2e-16 ***  |
| smokban          | 8.3339   | 10.8001    | 0.772   | 0.44703      |
| year1970         | -1.7684  | 0.2138     | -8.270  | 7.05e-09 *** |
| smokban:year1970 | -1.5008  | 0.4775     | -3.143  | 0.00403 **   |
| ---              |          |            |         |              |
- Signif. codes: 0 ‘\*\*\*’ 0.001 ‘\*\*’ 0.01 ‘\*’ 0.05 ‘.’ 0.1 ‘ ’ 1
- Residual standard error: 5.105 on 27 degrees of freedom
- Multiple R-squared: 0.9739, Adjusted R-squared: 0.971
- F-statistic: 335.7 on 3 and 27 DF, p-value: < 2.2e-16

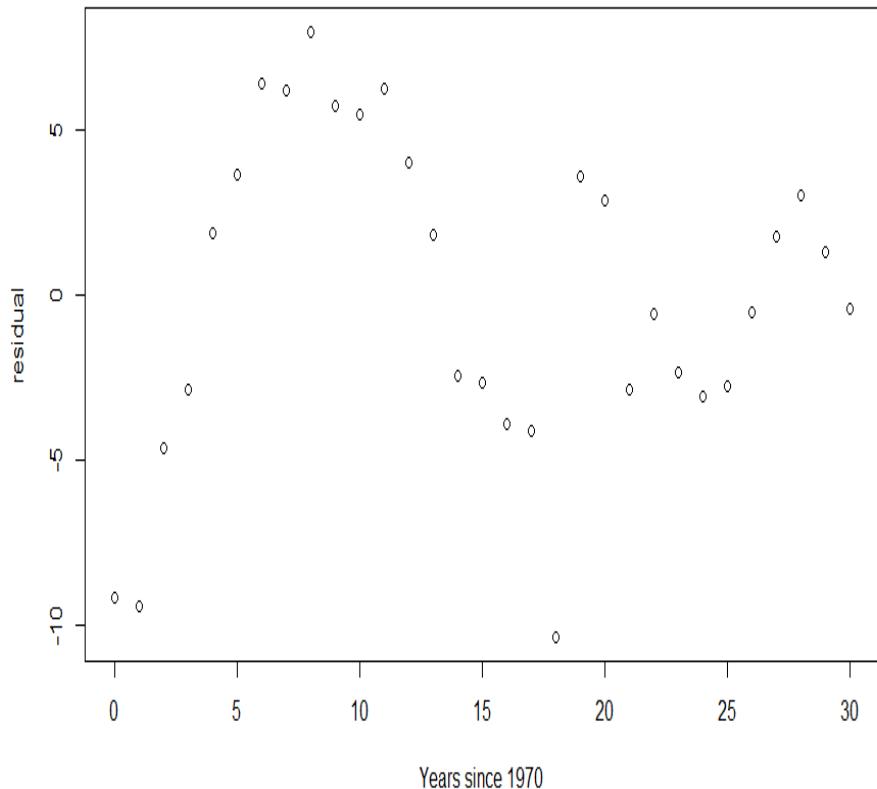


### California, 1970 - 2000





- `dwtest(cigsale.model01)`
- Durbin-Watson test
- data: `cigsale.model01`
- DW = 0.55595, p-value =  $8.806e-09$
- alternative hypothesis: true autocorrelation is greater than 0



# Model AR(1) directly

- Generalized least squares fit by REML
- Model: cigsale ~ smokban \* year1970
- Correlation Structure: AR(1). Formula: ~1
- Parameter estimate(s): Phi = 0.999993
- Coefficients:

|                  | Value     | Std.Error | t-value    | p-value |
|------------------|-----------|-----------|------------|---------|
| (Intercept)      | 123.00058 | 719.2730  | 0.1710068  | 0.8655  |
| smokban          | 28.09091  | 19.3476   | 1.4519066  | 0.1580  |
| year1970         | -1.83333  | 0.6335    | -2.8940451 | 0.0074  |
| smokban:year1970 | -1.80303  | 1.0286    | -1.7528736 | 0.0910  |
- Residual standard error: 719.273
- Degrees of freedom: 31 total; 27 residual



# Issues

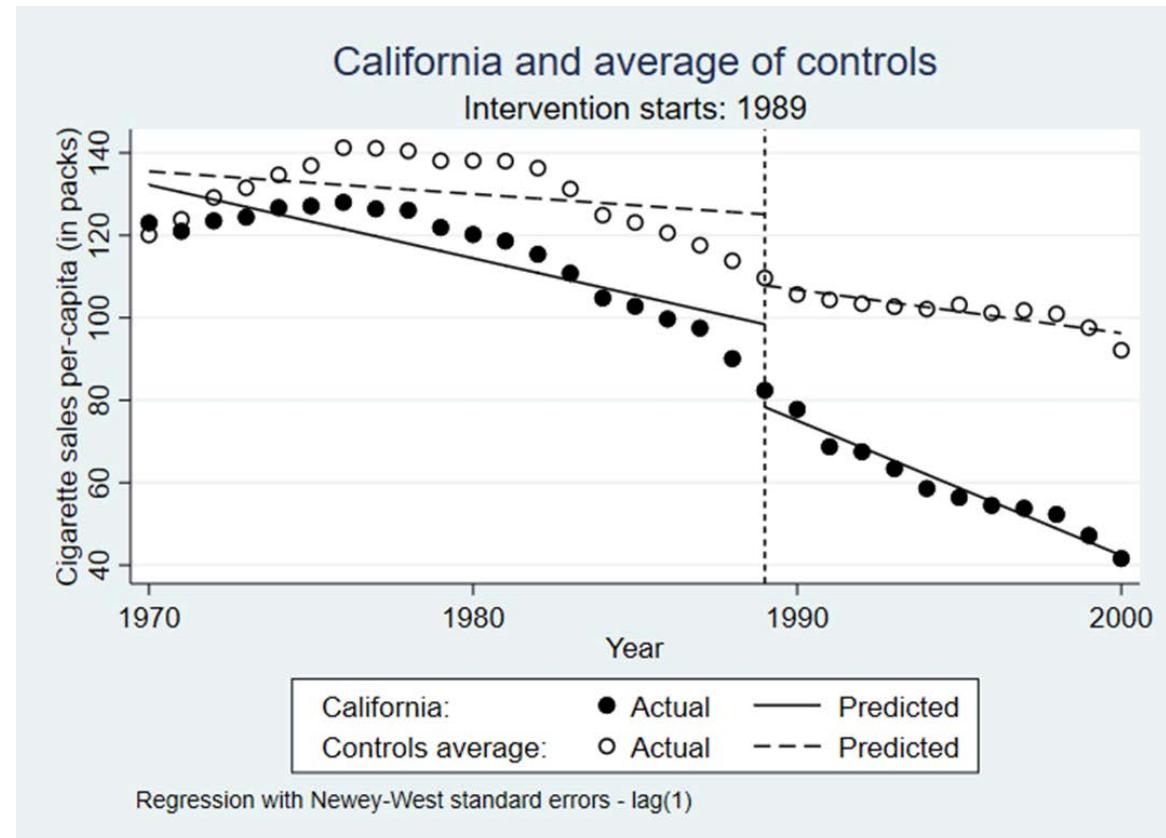
- How many observations?
  - Three?
  - Eight?





# Extensions

- Multiple-group design
- Compare intervention group with control group(s)

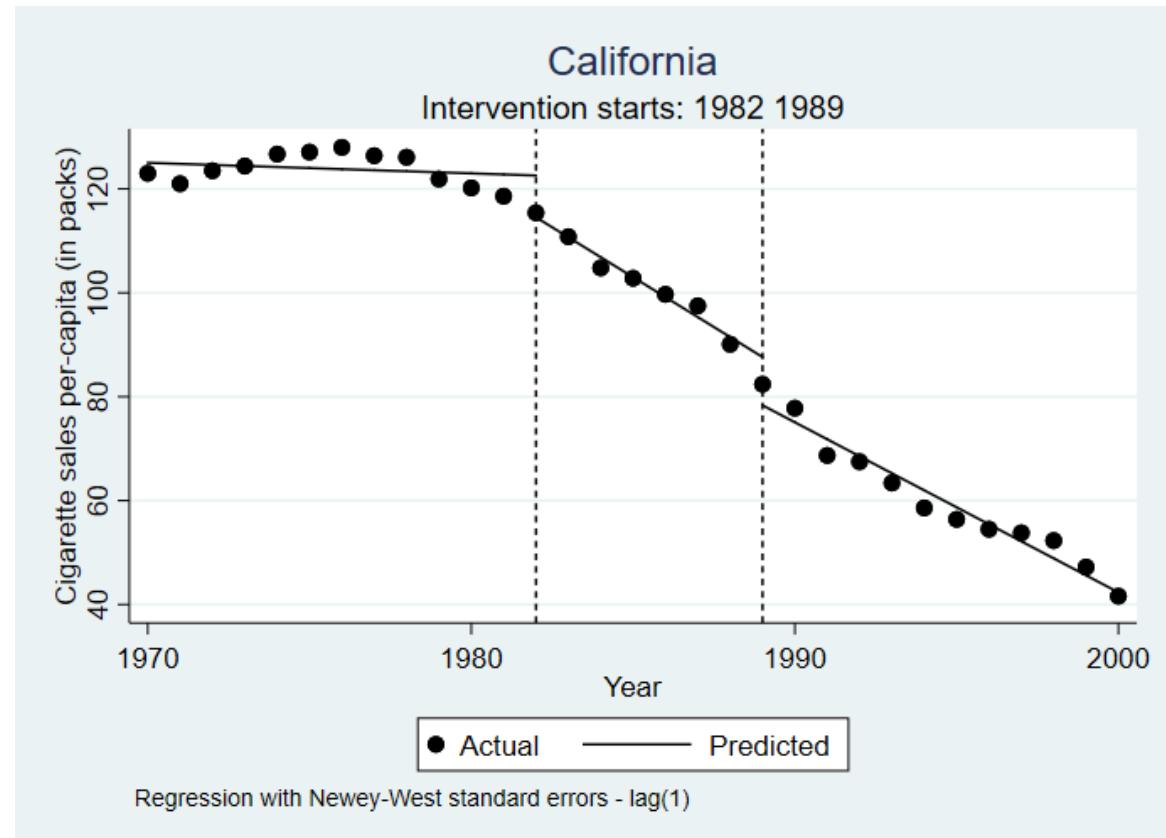




- **Types of control group**
  - Location e.g. state
  - Characteristic e.g. gender
  - Behaviour e.g. age group
  - Historical e.g. time periods
  - Outcome e.g. injury vs mortality
  - Time-period e.g. time of day



- Intervention starts and stops; or
- Multiple interventions

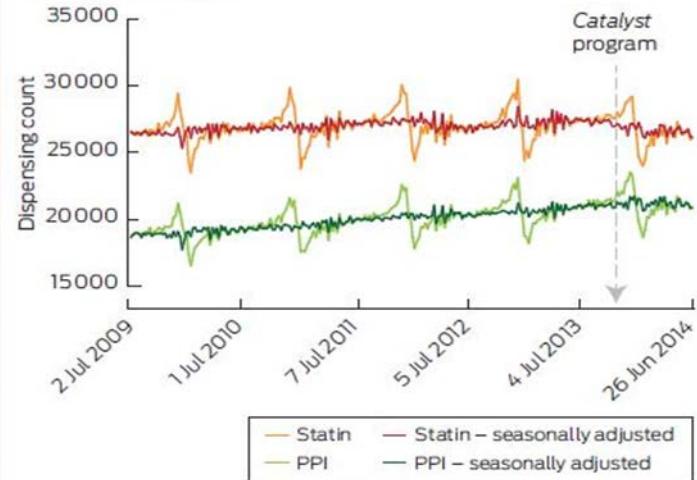




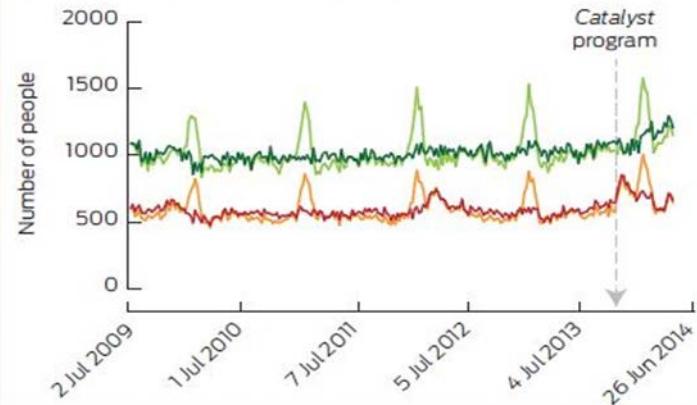
# • Seasonality

1 Weekly unadjusted and seasonally adjusted (A) dispensing counts and (B) number of people discontinuing use of statins and proton pump inhibitors (PPIs), 1 July 2009 to 30 June 2014

A: Dispensing counts



B: Number of people discontinuing use





- Extra variability
  - Increasing variance?
  - Count outcomes with Poisson distribution?
  - Model directly, not using itsa
- Time-varying confounders?
  - Include as covariates in regression model

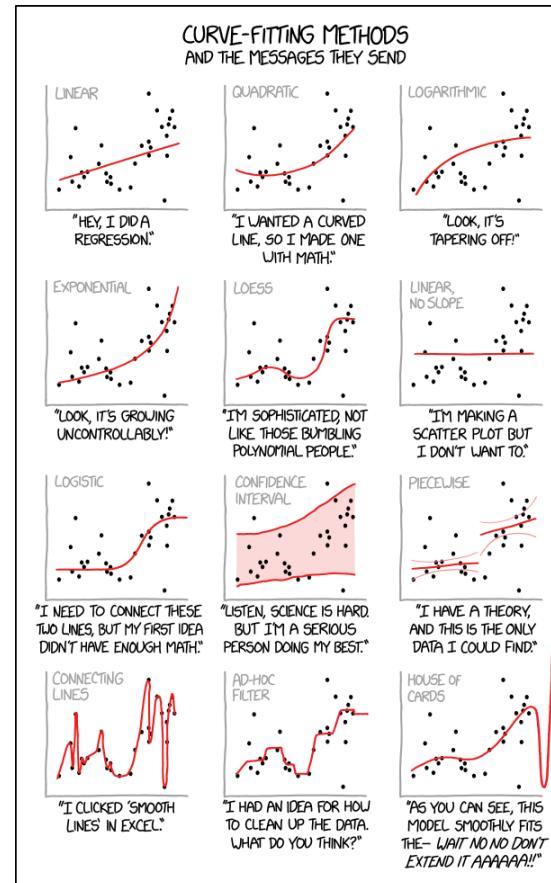


# Resources

- IDRE: nothing I could find
- Alternative names include changepoint analysis and piecewise or segmented regression



# Parting shot ...





# References

- Bernal et al (2017). Interrupted time series regression for the evaluation of public health interventions: a tutorial. *IJE* 46, 348 – 355.
- Bernal et al (2018). The use of controls in interrupted time series studies of public health interventions. *IJE* 47, 2082 – 2093.
- Bernal et al (2018). A methodological framework for model selection in interrupted time series studies. *J Clin Epi* 103, 82 – 91.
- Biglan et al. (2000). The value of interrupted time-series experiments for community intervention research. *Prevention Science* 1, 31-49.
- Schaffer et al. (2015). The crux of the matter: did the ABC's Catalyst program change statin use in Australia? *MJA* 202, 591 – 595.
- Linden (2015). Conducting interrupted time-series analysis for single- and multiple-group comparisons. *Stat Journal* 15, 480 – 500.



# Upcoming R workshops

10am-12, Tony McMichael seminar room

| Date         | Topic                             |
|--------------|-----------------------------------|
| 21 August    | Drop-in R setup                   |
| 22 August    | Packages, importing data, scripts |
| 5 September  | Visualisation (base R)            |
| 19 September | Visualisation (ggplot)            |
| 26 September | Meta analysis (for EPP)           |
| 3 October    | R Markdown                        |
| 17 October   | Logistic regression modelling     |
| 31 October   | Power and sample size             |